$ST/$ 02 Week 20 Simple Linear Regression Problem settings Given paired observations  $f(x_i, y_i)$   $y_{i-1}$  from model  $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$ with  $E(\mathcal{E}_i)=0$ , Vietri and Var  $(\mathcal{E}_i)$  =  $\alpha^2>0$ , Vietri Also assume  $Cov(\epsilon_j, \epsilon_j)$ =0 for all  $i \neq j$ . Here we focus on "fixed design" (then what is a "random design"?)  $\left\{\frac{1}{2}, \frac{1}{2}, \$  $Parameters$  to understand:  $\beta$ ,  $\beta$ ,  $\alpha^2$ .  $Fact: D E(y_i) = \beta_0 + \beta_1 x_i$ ,  $Var(y_i) = \alpha^2$ , and all yi's are uncorrelated. 2) If  $\mathcal{E}_i \wedge N(0, \infty^2) = 5$   $\mathcal{Y}_i \wedge N(\beta s \neq \beta, x_i, 0^2)$ . and Yi's are independent. LSE of parameters Define the lass function:  $L(\beta_{0},\beta_{1})=\frac{1}{2}\sum_{i=1}^{n}\sum_{i=1}^{n}(\gamma_{i}-\beta_{0}-\beta_{i}x_{i})^{2}$ 

Then let's find the minima of L:  $\frac{\partial}{\partial \beta_2}$  L( $\beta_0$ , $\beta_1$ ) = -2  $\frac{\sum_{j=1}^{n}(y_j - \beta_0 - \beta_1 x_j)$  (1)  $\frac{\partial}{\partial \beta_1}$  L( $\beta_0$ ,  $\beta_1$ ) = -2  $\sum_{i=1}^{n}$  X; ( $y_i$ ,  $\beta_0 - \beta_1$  X; ) (z)  $f(t)=0$  =>  $\beta = \bar{y} - \beta \bar{x}$ <br>(2) = 0  $\beta_1 = \frac{Z_7}{Z_1} (\frac{x_7 - \overline{x}}{y_1 - \overline{y}}) (\frac{y_7 - \overline{y}}{y_1} - \frac{\sum_{i=1}^{7} x_i y_i - n \overline{x} \overline{y}}{x_1^5 - \overline{x}^2 - n \overline{x}^2})$  $\sum_{j=1}^{z}$   $(x_j - \overline{x})^2$   $\sum_{j=1}^{z} x_j^2 - n\overline{x}$ What about the estimator of  $\alpha^2$ :  $\frac{2}{\pi}$  (g<sub>i</sub> -  $\beta$  -  $\beta$ , x, )<sup>2</sup><br> $n - 2$ Properties of estimators  $1) E(\beta - ) = \beta_0$  $Var(\hat{\beta}_{\circ}) = \frac{\sigma^2}{n} \sum_{\substack{i=1 \ i \neq j}}^{n} x_i^2$ 2)  $E(\hat{\beta}) = \beta$  $Var(\hat{\beta}_1) = \frac{\sigma^2}{\sum (x_i - \overline{x})^2}$ 

Inference for parameters in the case of Normal Further assume:  $\frac{\varepsilon_i}{i \cdot id}$ .  $3 = 3$  Yi $\sim N(s$ <sup>2</sup>  $s$ ,  $\lambda$ i,  $\sigma$ <sup>2</sup>) and *Yi's independent*. In addition,  $\beta_o \sim N(\beta_o, \frac{\alpha^2}{n} \frac{\sum_{i=1}^{N} x_i^2}{\sum_{i=1}^{N} (x_i - \overline{x})^2})$  $\beta$ i ~  $N$  $\beta$  $\frac{1}{2}$  $w_i$  th  $\alpha^2 = \frac{\sum_{i=1}^{n} (y_i - \beta_0 - \beta_1 x_i)^2}{n}$  $n - 2$ The estimated standard errors E.S.E.  $(\beta_0) = \frac{\partial}{\sqrt{n}} \left( \frac{\sum_{i=1}^{n} x_i^2}{\sum_{i=1}^{n} (x_i - \overline{x})^2} \right)^{\frac{1}{2}}$  $E.S.E. (\hat{\beta}) = \frac{\hat{\sigma}}{\int \hat{Z} \cdot (x)}$  $(X; - \times)$ <sup>2</sup>  $d_{emmn}$  1)  $\frac{(n-2)\sigma}{\sigma^2}$   $\sim$   $\pi^{-2}$ 2)  $\beta$  11  $\alpha$ <sup>2</sup>, and  $\beta$  -  $\beta$  1  $\alpha$  tn-2  $E$ .S.E.I $\beta$ b.

(continued) 3)  $\beta I \perp 0^{-4}$ , and  $\frac{\beta' - \beta}{2}$  $E.S.E.$   $\cancel{(}^\mathcal{B}$  $\mathcal{L}n$ -a Confidence Intervals  $(1 - d) \times 100%$  confidence interval for  $\beta$ o is  $\beta_0$  +  $t_{\frac{d}{2}, n-2}$   $E.S.E.\beta_0$ <sup>1</sup> <sup>d</sup> <sup>100</sup> confidence interval for β is  $\hat{\beta}$ , <u>t</u> t<u>o</u>, n-2 E.S.E. (β) Test the slope  $H_0$   $\rho_1 = b$   $v.s.$   $H_1$  : ...  $\mu = \frac{\beta_1 - \beta_2}{\beta_2 - \beta_1}$ <br>E.S.E.( $\beta$  $\beta$ i tn-2 under Ho Verification by MLE (still assume Normal)  $L(\beta_0, \beta_1, \alpha^2) = \frac{R}{i^2} \frac{1}{\sqrt{2\pi\alpha^2}} \exp\left\{-\frac{1}{2\alpha^2} (y_i - \beta_0 - \beta_1 x_i)^2\right\}$  $f(x) = \left(\frac{1}{x^{2}}\right)^{\frac{n}{2}} exp\left(-\frac{1}{2x^{2}}\sum_{i=1}^{n}(y_{i}-\beta_{i}-\beta_{i}x_{i})^{2}\right)$ 

(continued) Then we can further have the log-likelihood:  $l$  (β<sub>D</sub>, β<sub>1</sub>, 0<sup>2</sup>) = <u>n</u>  $ln(\frac{1}{0^{2}})-\frac{1}{10^{2}}\sum_{i=1}^{n}(y_{i}-\beta_{0}-\beta_{i}x_{i})^{2}$ +c Which part is flexible? How to maximize it? (C) Tao Ma All Rights Reserved